

Time-Varying Expected Consumption Growth and Cross-Sectional Equity Returns

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Outline

- Introduction and Literature Review
- Consumption-based Two-beta Model
- Empirical Specification and Methods
- Results
- Robustness Check
- Conclusions and Future Work

Motivation

- Stock returns exhibit cross-sectional predictive patterns (ME, B/M, E/P, CF/P...). Behavior bias or risks?
- Traditional CCAPM testing typically assumes i.i.d consumption growth and works poorly in explaining the cross-sectional patterns.
- Early evidence shows predictability in economic and consumption growth (e.g. Harvey(1988), Kandel/Stambaugh(1990)).

Goals:

- A risk-based explanation for the cross-sectional return patterns, particularly size and B/M effects.
- A structural link between macroeconomic variables and systematic risks in the financial market through predictability of aggregate cash flow
- Cash-flow perspective for studying intertemporal asset-pricing models

Main Results:

- Non-expected utility together with a VAR assumption on the aggregate consumption growth dynamics leads to an intertemporal asset-pricing model where an asset's expected return is determined by its covariances with two types of consumption growth shocks: the unexpected consumption growth (*"current growth shock"*) and the shock to the expected future growths (*"expected growth shock"*). This leads to a two-beta model for the explaining the cross-sectional asset returns.
- There is strong empirical support for long-term predictability in the aggregate consumption growth;
- The news on *expected consumption growth* contains valuable information for explaining a wide variety of cross-sectional return patterns.

Related Literature

Literature	Comparison
Aggregate cash flow dynamics has asset-pricing implications: Bansal/Yaron (2003) Bansal/Gallant/Tauchen (2002) Tauchen (2004)	Empirically identify expected growth news; cross-sectional return puzzles; link to macroeconomic variables
Macroeconomic variables as risk factors: Chen/Roll/Ross (1986) Chen (1991) Ferson/Harvey (1999)	Equilibrium restrictions from growth predictability
ICAPM: Campbell(1996) Campbell/Vuolteenaho(2003)	Cash flow perspective

Two Growth Shocks

Assume a market portfolio which pays dividend equal to the consumption and demands return $R_{a,t+1}$.

$$R_{a,t+1} = \frac{P_{t+1} + C_{t+1}}{P_t} \quad (1)$$

It can be shown that

$$\begin{aligned} r_{a,t+1} - E_t[r_{a,t+1}] &= g_{c,t+1} - E_t[g_{c,t+1}] + (E_{t+1} - E_t) \left[\sum_{j=1}^{\infty} \kappa_{c,1}^j g_{c,t+1+j} \right] \\ &\quad - (E_{t+1} - E_t) \left[\sum_{j=1}^{\infty} \kappa_{c,1}^j r_{a,t+1+j} \right] \end{aligned} \quad (2)$$

Two Growth Shocks

Two questions:

1. Why is it important to differentiate the two types of shocks?
2. How to identify the two types shocks?

Pricing Kernel

A representative investor maximizes his utility which is assumed to take the recursive form as in Epstein and Zin (1989). That is,

$$U(C_t, E_t(U_{t+1})) = \left[(1 - \delta)C_t^{\frac{1-\gamma}{\theta}} + \delta(E_t U_{t+1}^{1-\gamma})^{\frac{1}{\theta}} \right]^{\frac{\theta}{1-\gamma}} \quad (3)$$

where $\theta \equiv \frac{1-\gamma}{1-\frac{1}{\psi}}$. $\gamma > 0$ is the risk-aversion coefficient and $\psi > 0$ is the inter-temporal elasticity of substitution(IES). If $\theta = 1$ or $\gamma = \frac{1}{\psi}$, then (3) is time separable.

From the Euler equation, we can get the log pricing kernel to be

$$m_{t+1} = \text{const} - \frac{\theta}{\psi} g_{c,t+1} + (\theta - 1)r_{a,t+1} \quad (4)$$

Applying $E_t[m_{t+1}r_{a,t+1}] = 0$ and assuming joint homoscedasticity for $g_{c,t+1}$ and $r_{a,t+1}$, we get

$$E_t g_{c,t+1} = \mu_m + \psi E_t r_{a,t+1} \quad (5)$$

Substitute into (2), we get

$$\underbrace{r_{a,t+1} - E_t[r_{a,t+1}]}_{\eta_{r,t+1}} = \underbrace{g_{c,t+1} - E_t[g_{c,t+1}]}_{\eta_{c,t+1}} + \underbrace{\left(1 - \frac{1}{\psi}\right) (E_{t+1} - E_t) \left[\sum_{j=1}^{\infty} \kappa_{c,1}^j g_{c,t+1+j} \right]}_{\epsilon_{c,t+1}} \quad (6)$$

Write the pricing kernel in terms of innovations using (6), we have

$$\eta_{m,t+1} = \left[-\frac{\theta}{\psi} + (\theta - 1)\right] \eta_{c,t+1} + (\theta - 1) \left(1 - \frac{1}{\psi}\right) \epsilon_{c,t+1} \quad (7)$$

VAR

The consumption growth is modelled as following a VAR process as

$$\begin{bmatrix} g_{c,t+1} \\ x_{t+1} \end{bmatrix} = \begin{bmatrix} a_{11} & a_{12} \\ a_{21} & a_{22} \end{bmatrix} \begin{bmatrix} g_{c,t} \\ x_t \end{bmatrix} + \begin{bmatrix} \eta_{c,t+1} \\ \eta_{x,t+1} \end{bmatrix} \quad (8)$$

or

$$Z_{t+1} = AZ_t + \omega_{t+1}$$

Define a $K \times 1$ vector e_1 which has first element being 1 and the rest being 0. Hence $e_1'Z_t$ picks out the consumption growth from the vector Z_t . We can then get

$$\eta_{c,t+1} = \omega_{1,t+1} \quad (9)$$

$$\epsilon_{c,t+1} = e_1' \kappa_{1,c} A (I - \kappa_{1,c} A)^{-1} \omega_{t+1} = \delta' \omega_{t+1} \quad (10)$$

Two-Beta Model

The following equation needs to hold for any asset in the economy.

$$E_t[r_{i,t+1} - r_{f,t+1}] + \frac{\sigma_{i,t}^2}{2} = -Cov_t(m_{t+1}, r_{i,t+1}) \quad (11)$$

$$E_t[r_{i,t+1} - r_{f,t+1}] + \frac{\sigma_{i,t}^2}{2} = \gamma Cov_t(\eta_{c,t+1}, r_{i,t+1}) + \left(\gamma - \frac{1}{\psi}\right) Cov_t(\epsilon_{c,t+1}, r_{i,t+1}) \quad (12)$$

$$= \gamma \sigma_{\eta}^2 \beta_{i,\eta} + \left(\gamma - \frac{1}{\psi}\right) \sigma_{\epsilon}^2 \beta_{i,\epsilon} \quad (13)$$

where

$$\beta_{i,\eta} = \frac{Cov(\eta_{c,t+1}, r_{i,t+1})}{\sigma_{\eta}^2} \quad (14)$$

$$\beta_{i,\epsilon} = \frac{Cov(\epsilon_{c,t+1}, r_{i,t+1})}{\sigma_{\epsilon}^2} \quad (15)$$

Empirical Methods

1. Estimate the VAR system and obtain the time-series of $\eta_{c,t}$ and $\epsilon_{c,t}$.
2. Run time-series regression of each portfolio return series separately on $\eta_{c,t}$ and $\epsilon_{c,t}$ to obtain $\beta_{i,\eta}$ and $\beta_{i,\epsilon}$.

$$r_{i,t} = \alpha_i + \beta_{i,\eta}\eta_{c,t} + u_{i,t}$$

$$r_{i,t} = \phi_i + \beta_{i,\epsilon}\epsilon_{c,t} + \nu_{i,t}$$

3. Run cross-sectional regression of average return on the betas.

$$\bar{r}_i = \lambda_0 + \lambda_1\beta_{i,\eta} + e_i$$

$$\bar{r}_i = \lambda_0 + \lambda_2\beta_{i,\epsilon} + e_i$$

$$\bar{r}_i = \lambda_0 + \lambda_1\beta_{i,\eta} + \lambda_2\beta_{i,\epsilon} + e_i$$

GMM Moment conditions for step 2-3

$$E[u_{i,t}] = 0;$$

$$E[u_{i,t}\eta_{c,t}] = 0;$$

$$E[\nu_{i,t}] = 0;$$

$$E[\nu_{i,t}\epsilon_{c,t}] = 0;$$

$$\sum_{i=1}^N E\{r_{i,t} - \lambda_0 - \lambda_1\beta_{\eta,i} - \lambda_2\beta_{\epsilon,i}\} = 0;$$

$$\sum_{i=1}^N E\{[r_{i,t} - \lambda_0 - \lambda_1\beta_{\eta,i} - \lambda_2\beta_{\epsilon,i}]\beta_{\eta,i}\} = 0;$$

$$\sum_{i=1}^N E\{[r_{i,t} - \lambda_0 - \lambda_1\beta_{\eta,i} - \lambda_2\beta_{\epsilon,i}]\beta_{\epsilon,i}\} = 0;$$

Data

VAR Data

- $g_{c,t}$, the growth of real per capita consumption;
- $Term_t$, the yield spread between 10-year T-bill and 3-month T-bill
- Div_t , the deseasonalized dividend-price ratio for the CRSP stock index
- Nrf_t , the quarterly risk free rate measured by the one-month T-bill rates compounded over the quarter
- $Infl_t$, realized inflation rate in period t from PCE index
- Rrf_t , quarter risk free rate = $(1 + Nrf_t)/(1 + Infl) - 1$
- Def_t , the yield spread between the Moody's rated Baa bonds and Aaa bonds

Return Data

- $r_{m,t}$, the CRSP value-weighted index return in excess of the risk-free rate
- 25 Size-B/M double-sorted portfolios;
- Size, B/M, Earning/Price, Cash-flow/Price sorted deciles;

Table 1: Consumption Growth Prediction: Univariate

1993.1–2002.4

Variable	$q1$	$q2$	$q3$	$q4$	$q5$	$q6$	$q7$	$q8$
$Term_t$	0.085 (0.032)	0.160 (0.052)	0.233 (0.065)	0.286 (0.080)	0.314 (0.093)	0.321 (0.102)	0.313 (0.112)	0.288 (0.125)
Div_t	-0.195 (0.132)	-0.250 (0.214)	-0.275 (0.275)	-0.310 (0.351)	-0.299 (0.420)	-0.302 (0.481)	-0.323 (0.538)	-0.325 (0.604)
Rrf_t	-0.008 (0.093)	0.002 (0.165)	0.050 (0.233)	0.110 (0.282)	0.155 (0.319)	0.248 (0.357)	0.361 (0.391)	0.460 (0.438)
$Infl_t$	-0.212 (0.062)	-0.406 (0.117)	-0.607 (0.174)	-0.771 (0.240)	-0.886 (0.298)	-1.011 (0.354)	-1.108 (0.411)	-1.151 (0.467)
Def_t	-0.055 (0.128)	-0.073 (0.253)	-0.089 (0.377)	-0.027 (0.481)	0.018 (0.570)	0.086 (0.644)	0.181 (0.704)	0.243 (0.760)

Table 2: Consumption Growth Prediction: Multivariate

Horizon	Intercept	$Term_t$	Div_t	Rrf_t	$Infl_t$	$Adj. R^2$
$q1$	0.225 (0.091)	0.012 (0.041)	0.554 (0.234)	-0.218 (0.090)	-0.443 (0.116)	0.148 .
$q2$	0.475 (0.167)	0.003 (0.074)	1.368 (0.436)	-0.463 (0.166)	-0.970 (0.192)	0.234 .
$q3$	0.695 (0.230)	-0.008 (0.103)	2.222 (0.604)	-0.677 (0.229)	-1.503 (0.253)	0.305 .
$q4$	0.840 (0.303)	-0.019 (0.134)	2.826 (0.727)	-0.816 (0.301)	-1.902 (0.317)	0.307 .
$q5$	0.978 (0.401)	-0.046 (0.165)	3.398 (0.883)	-0.949 (0.398)	-2.249 (0.385)	0.300 .
$q6$	1.063 (0.500)	-0.092 (0.193)	3.934 (1.025)	-1.028 (0.496)	-2.586 (0.454)	0.300 .
$q7$	1.069 (0.600)	-0.133 (0.222)	4.258 (1.115)	-1.028 (0.595)	-2.798 (0.547)	0.279 .
$q8$	1.026 (0.712)	-0.184 (0.257)	4.470 (1.319)	-0.980 (0.707)	-2.909 (0.664)	0.246 .

Figure 1: Time-Variation in Expected Consumption Growth

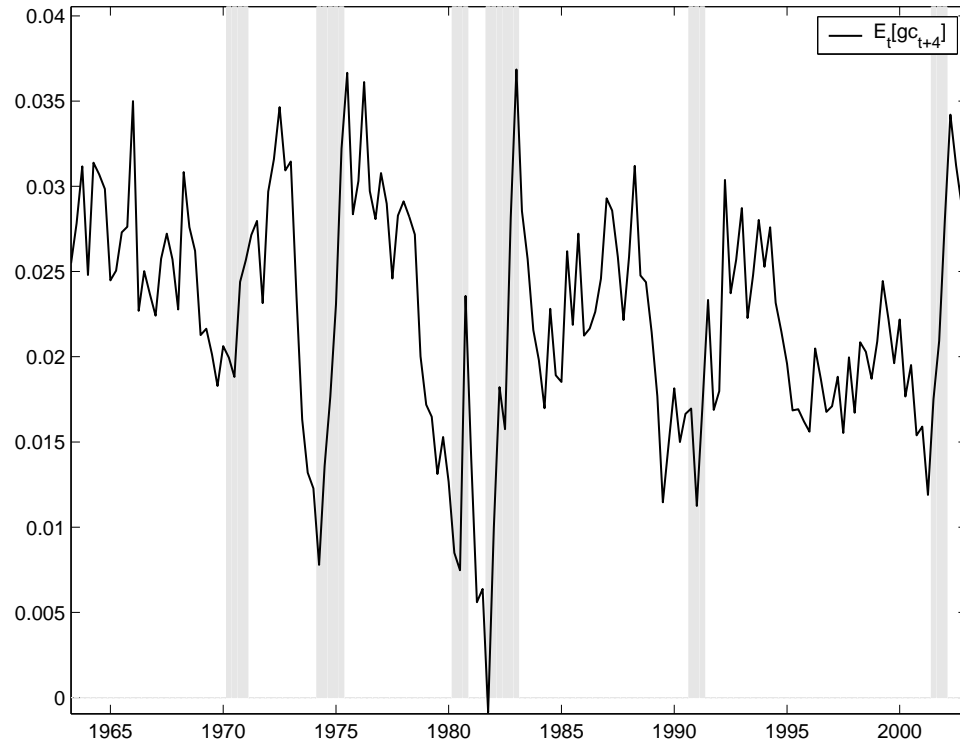


Table 3: VAR Estimations

	$g_{c,t}$	$Term_t$	Div_t	Rrf_t	$Infl_t$	$Adj. R^2$
$g_{c,t+1}$	0.299 (0.061)	0.022 (0.038)	0.392 (0.179)	-0.181 (0.069)	-0.311 (0.088)	0.229
$Term_{t+1}$	-0.355 (0.117)	0.875 (0.047)	0.373 (0.296)	0.184 (0.113)	-0.144 (0.147)	0.785
Div_{t+1}	0.028 (0.014)	-0.003 (0.004)	0.884 (0.054)	0.017 (0.015)	0.053 (0.024)	0.925
Rrf_{t+1}	-0.001 (0.065)	0.038 (0.030)	-0.334 (0.174)	0.890 (0.056)	0.223 (0.113)	0.610
$Infl_{t+1}$	0.153 (0.050)	-0.044 (0.027)	0.212 (0.171)	0.009 (0.051)	0.837 (0.099)	0.795

VAR Implied Long-Run Statistics

$$Z_{t+1} = AZ_t + \omega_{t+1}$$

$$V = E(\omega_{t+1}\omega'_{t+1})$$

$$C(0) = \sum_{j=0}^{\infty} A^j V A^{j'}; \quad C(j) = A^j C(0)$$

$$V_k = kC(0) + \sum_{j=1}^{k-1} (k-j)[C(j) + C(j)'];$$

$$W_k = \sum_{j=1}^k (I - A)^{-1} (I - A^j) V (I - A^j)' (I - A)^{-1'}$$

$$R^2(k) = 1 - \frac{e1' W_k e1}{e1' V_k e1}$$

$$VR(k) = \frac{e1' V_k e1}{k e1' C(0) e1}$$

Figure 2: Long-run Consumption Growth Predictability Implied by VAR

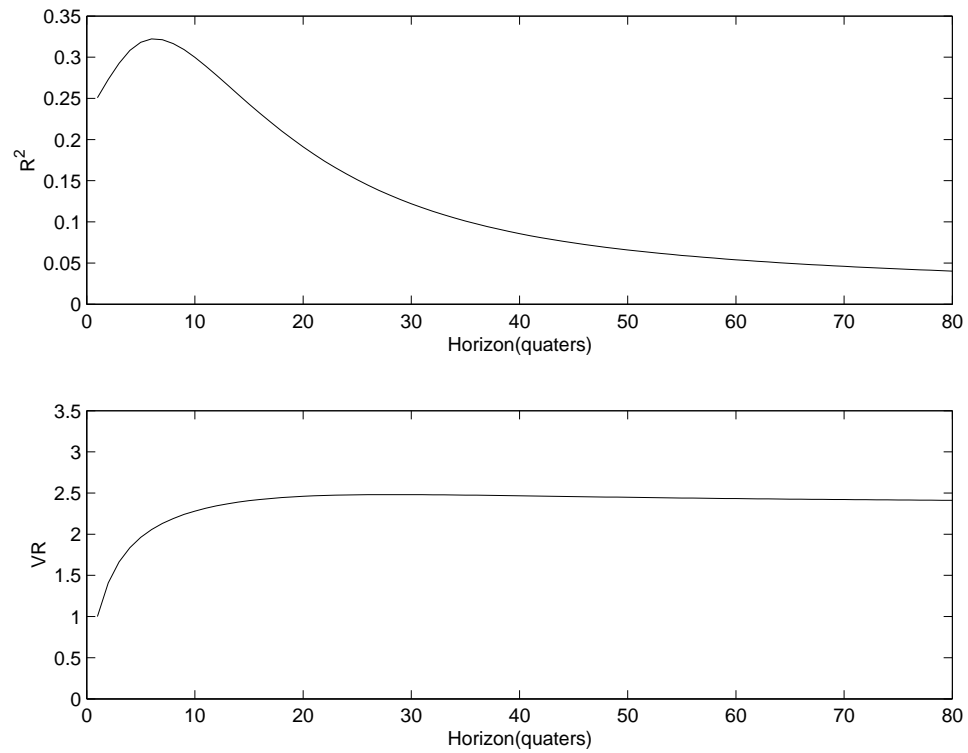
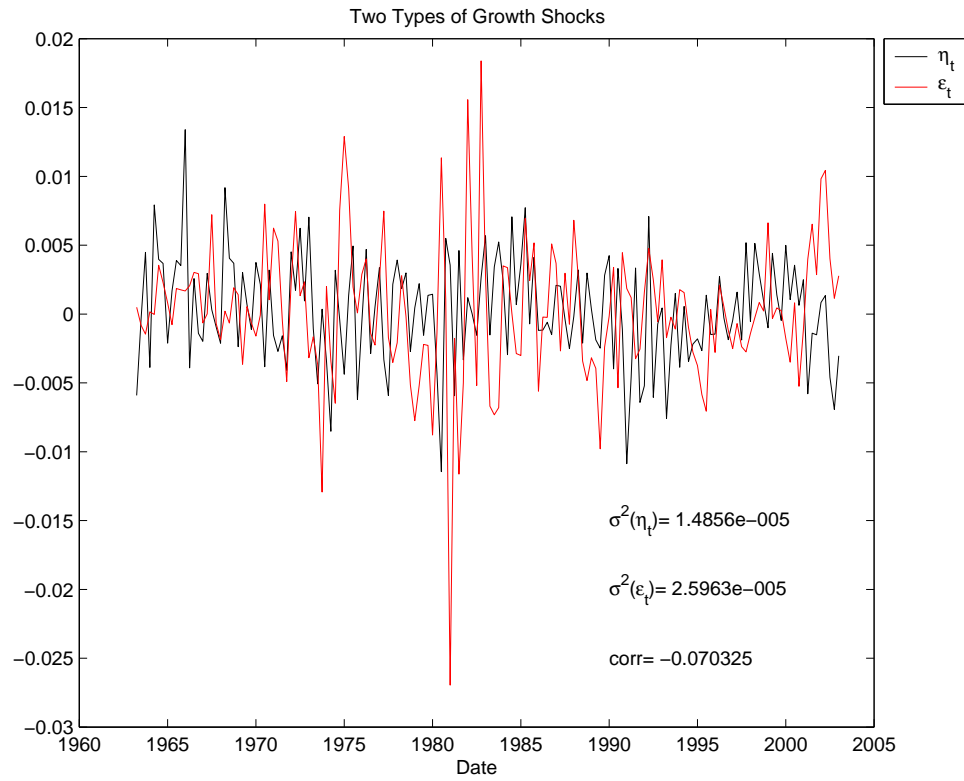


Figure 3: Time-Series of Two Shocks



$$\eta_{c,t} = \omega_{1,t}$$

$$\epsilon_{c,t} = e_1' \kappa_c A (I - \kappa_c A)^{-1} \omega_t$$

Table 4: Exposure to Fama-French 3 Factors

	ME1	2	3	4	ME5	ME1	2	3	4	ME5
	Average Excess Returns					s.e.				
BM1	0.931	1.196	1.251	1.582	1.308	0.031	0.024	0.022	0.021	0.020
2	2.466	2.025	2.137	1.477	1.326	0.028	0.022	0.022	0.018	0.017
3	2.631	2.723	2.166	2.122	1.433	0.027	0.024	0.021	0.019	0.016
4	3.326	2.948	2.607	2.540	1.720	0.028	0.025	0.022	0.021	0.017
BM5	3.637	3.153	3.057	2.721	1.762	0.031	0.026	0.025	0.022	0.020
	β_{MKT}					s.e.				
1	1.038	1.075	1.036	1.031	1.016	0.049	0.032	0.027	0.038	0.023
2	0.978	1.000	1.033	1.052	1.021	0.035	0.025	0.033	0.024	0.034
3	0.924	1.003	0.996	1.050	0.926	0.039	0.026	0.030	0.023	0.038
4	0.904	1.021	1.046	1.078	1.005	0.047	0.023	0.035	0.041	0.030
5	1.023	1.080	1.055	1.100	1.043	0.037	0.038	0.036	0.050	0.038
	β_{SMB}					s.e.				
1	1.455	1.028	0.729	0.369	-0.270	0.068	0.044	0.044	0.066	0.040
2	1.336	0.969	0.584	0.312	-0.225	0.052	0.053	0.042	0.073	0.050
3	1.159	0.764	0.499	0.218	-0.232	0.054	0.041	0.041	0.047	0.050
4	1.112	0.702	0.418	0.202	-0.190	0.059	0.034	0.049	0.046	0.034
5	1.195	0.814	0.616	0.407	-0.122	0.049	0.050	0.051	0.066	0.072
	β_{HML}					s.e.				
1	-0.411	-0.444	-0.509	-0.504	-0.314	0.085	0.060	0.043	0.050	0.052
2	0.075	0.139	0.165	0.230	0.124	0.048	0.096	0.105	0.132	0.107
3	0.307	0.377	0.467	0.468	0.281	0.051	0.094	0.112	0.108	0.081
4	0.437	0.614	0.694	0.596	0.586	0.039	0.089	0.113	0.080	0.092
5	0.730	0.808	0.817	0.746	0.641	0.048	0.048	0.059	0.076	0.063

Table 5: Exposure to the two types of consumption risks for FF25

	ME1	2	3	4	ME5	ME1	2	3	4	ME5
	Average Excess Returns					s.e.				
BM1	0.931	1.196	1.251	1.582	1.308	0.031	0.024	0.022	0.021	0.020
2	2.466	2.025	2.137	1.477	1.326	0.028	0.022	0.022	0.018	0.017
3	2.631	2.723	2.166	2.122	1.433	0.027	0.024	0.021	0.019	0.016
4	3.326	2.948	2.607	2.540	1.720	0.028	0.025	0.022	0.021	0.017
BM5	3.637	3.153	3.057	2.721	1.762	0.031	0.026	0.025	0.022	0.020
	β_η					s.e.				
1	8.598	6.684	5.378	5.187	3.811	2.769	2.761	2.737	2.700	2.164
2	8.516	5.346	4.634	4.210	2.096	2.322	2.137	2.075	1.969	1.767
3	6.223	4.839	4.168	3.366	2.558	2.144	2.088	1.957	2.015	1.496
4	6.017	4.367	3.543	2.714	1.870	2.063	2.187	2.218	2.302	1.700
5	5.995	5.725	4.424	3.947	3.070	2.379	2.477	2.302	2.515	1.753
	β_ϵ					s.e.				
1	-1.844	-1.462	-0.721	-0.909	-1.198	2.593	2.402	1.992	1.888	1.547
2	-0.502	-0.583	-0.059	-0.983	-1.435	2.327	2.060	1.975	1.894	1.475
3	0.181	0.137	-0.024	-0.357	-1.568	2.041	1.896	1.788	1.859	1.428
4	0.361	0.248	0.345	0.211	-0.304	2.048	1.902	1.856	1.955	1.529
5	0.645	0.136	0.175	-0.051	-0.341	2.133	2.086	1.971	2.099	1.503

Table 6: Exposure to the Campbell(1996) Factors for FF25

	ME1	2	3	4	ME5	ME1	2	3	4	ME5
	Average Excess Returns					s.e.				
BM1	0.931	1.196	1.251	1.582	1.308	0.031	0.024	0.022	0.021	0.020
2	2.466	2.025	2.137	1.477	1.326	0.028	0.022	0.022	0.018	0.017
3	2.631	2.723	2.166	2.122	1.433	0.027	0.024	0.021	0.019	0.016
4	3.326	2.948	2.607	2.540	1.720	0.028	0.025	0.022	0.021	0.017
BM5	3.637	3.153	3.057	2.721	1.762	0.031	0.026	0.025	0.022	0.020
	β_a					s.e.				
1	1.656	1.576	1.456	1.350	1.041	0.110	0.097	0.084	0.080	0.038
2	1.411	1.280	1.168	1.092	0.918	0.089	0.086	0.068	0.079	0.048
3	1.213	1.126	1.004	0.977	0.755	0.092	0.090	0.087	0.070	0.042
4	1.140	1.040	0.940	0.944	0.749	0.093	0.094	0.093	0.076	0.068
5	1.163	1.073	0.991	0.981	0.784	0.105	0.107	0.103	0.105	0.080
	β_h					s.e.				
1	-1.418	-1.380	-1.274	-1.184	-0.958	0.130	0.124	0.103	0.087	0.063
2	-1.227	-1.159	-1.051	-1.022	-0.850	0.114	0.094	0.082	0.061	0.057
3	-1.100	-1.030	-0.928	-0.909	-0.698	0.090	0.084	0.079	0.061	0.060
4	-1.035	-0.959	-0.877	-0.875	-0.700	0.087	0.078	0.064	0.069	0.056
5	-1.044	-0.971	-0.902	-0.916	-0.723	0.089	0.086	0.091	0.084	0.074
	β_{CAPM}					s.e.				
1	1.647	1.556	1.442	1.318	1.032	0.071	0.066	0.058	0.059	0.022
2	1.387	1.269	1.168	1.076	0.907	0.061	0.063	0.052	0.066	0.044
3	1.199	1.126	1.003	0.965	0.757	0.081	0.071	0.074	0.060	0.043
4	1.119	1.044	0.951	0.945	0.748	0.085	0.074	0.077	0.064	0.058
5	1.168	1.076	0.983	0.984	0.789	0.084	0.081	0.086	0.083	0.063

Table 7: Cross-sectional Test Results

	λ_0	λ_η	λ_ϵ	λ_a	λ_h	λ_m	λ_g	λ_{smb}	λ_{hml}	$Adj. R^2$
β_η	1.957 (0.704)	0.045 (0.145)	-0.032
β_ϵ	2.557 (1.891)	.	0.977 (0.522)	0.811
$\beta_\eta \& \beta_\epsilon$	2.221 (1.131)	0.073 (0.141)	0.990 (0.258)	0.833
β_a	2.959 (1.054)	.	.	-0.709 (1.044)	0.009
β_h	2.993 (1.171)	.	.	.	0.817 (1.293)	0.001
$\beta_a \& \beta_h$	1.467 (1.190)	.	.	-10.596 (6.236)	-12.402 (7.338)	0.022
<i>CAPM</i>	2.943 (1.090)	-0.699 (1.085)	.	.	.	0.005
<i>CCAPM</i>	1.460 (0.880)	0.180 (0.198)	.	.	0.033
<i>FF3</i>	2.717 (1.585)	-1.308 (1.463)	.	0.680 (0.530)	1.444 (0.517)	0.764

Figure 4: Goodness of Fit

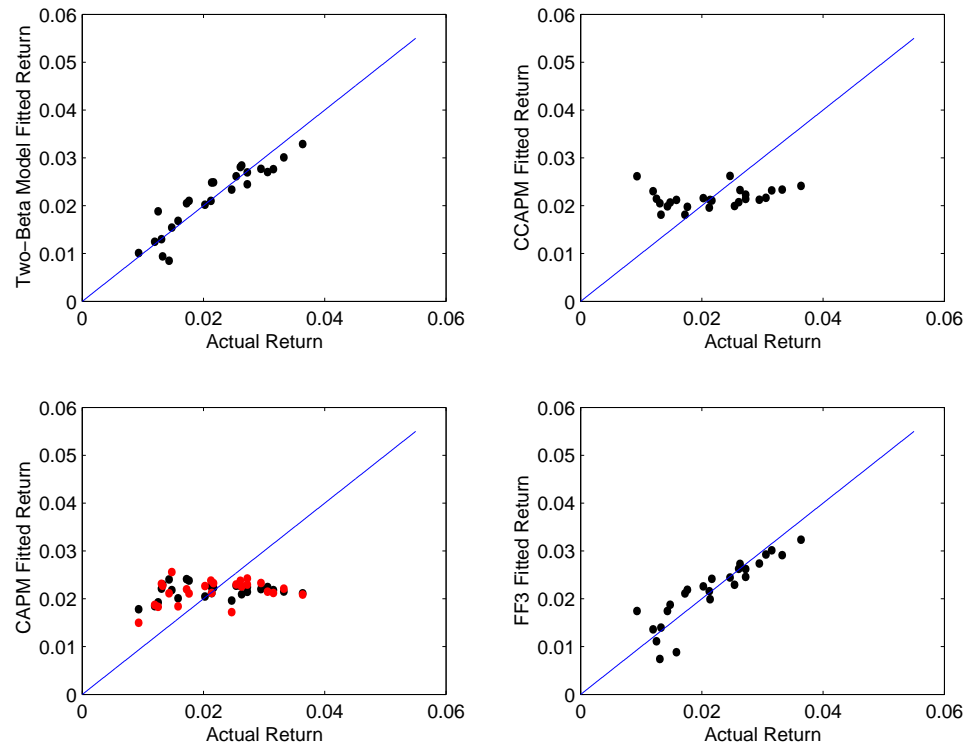


Table 8: Cross-sectional Results for Fama-French 25 Portfolios

	λ_0	λ_η	λ_ϵ	λ_{mkt}	λ_{smb}	λ_{hml}	$\log(ME)$	$\log(BM)$	$Adj. R^2$
$\beta_\eta, \beta_\epsilon, FF3$	2.707 (2.687)	0.140 (0.150)	0.663 (0.230)	-1.105 (3.013)	-0.028 (1.374)	0.747 (0.734)	.	.	0.883
<i>Size</i>	4.268 (0.832)	-0.249 (0.097)	.	0.187
<i>B/M</i>	2.251 (0.095)	0.861 (0.139)	0.607
<i>Size + B/M</i>	3.696 (0.513)	-0.172 (0.060)	0.790 (0.124)	0.700
$\beta_\eta, \beta_\epsilon, Size$	1.885 (3.183)	0.093 (0.236)	1.016 (0.442)	.	.	.	0.030 (0.214)	.	0.825
$\beta_\eta, \beta_\epsilon, B/M$	2.010 (1.390)	0.102 (0.175)	0.705 (0.269)	0.397 (0.386)	0.889
$\beta_\eta, \beta_\epsilon, Size, B/M$	0.503 (2.540)	0.195 (0.158)	0.786 (0.270)	.	.	.	0.131 (0.180)	0.446 (0.454)	0.895

Robustness Checks

- Other predictive variables in the VAR system.
- Different sets of portfolios;
- Sub-sample studies (for example, pre-1990 sample).

Table 9: Robustness to Choices of Information Variables: FF25 portfolios

Model	λ_0	λ_η	λ_ϵ	\bar{R}^2
<i>Term + Div + nrf + Def</i>	2.498 (0.910)	-0.003 (0.160)	0.872 (0.230)	0.783
<i>Term + Div + Rrf + Infl + Def</i>	3.080 (1.033)	0.423 (0.185)	0.979 (0.260)	0.739
<i>Term + Div + Rrf + Infl</i>	2.684 (0.973)	0.072 (0.147)	0.980 (0.265)	0.833
<i>Div + Infl + Term</i>	2.461 (0.962)	0.109 (0.132)	0.937 (0.246)	0.836
<i>Div + Rrf + Infl</i>	1.473 (0.969)	0.087 (0.155)	1.258 (0.377)	0.612
<i>Div + Rrf + Term</i>	2.261 (0.981)	-0.268 (0.136)	0.677 (0.173)	0.750
<i>Rrf + Infl + Term</i>	0.933 (1.072)	-0.205 (0.130)	0.912 (0.217)	0.530
<i>Div^{rep} + Rrf + Infl + Term</i>	1.828 (0.923)	-0.122 (0.133)	0.793 (0.205)	0.686

Table 10: Cross-sectional Results for Size,B/M, E/P, CF/P Portfolios

	λ_0	λ_η	λ_ϵ	λ_a	λ_h	λ_m	λ_g	λ_{smb}	λ_{hml}	$Adj. R^2$
β_η	1.335 (0.785)	0.134 (0.125)	0.084
β_ϵ	2.354 (1.289)	.	0.685 (0.336)	0.611
$\beta_\eta \& \beta_\epsilon$	1.893 (1.067)	0.120 (0.131)	0.674 (0.307)	0.692
β_a	2.108 (1.001)	.	.	-0.270 (0.956)	-0.019
β_h	2.368 (1.083)	.	.	.	0.579 (1.154)	-0.004
$\beta_a \& \beta_h$	3.715 (1.913)	.	.	12.824 (7.734)	16.048 (9.445)	0.228
<i>CAPM</i>	1.602 (1.030)	-0.178 (0.982)	.	.	.	-0.023
<i>CCAPM</i>	0.662 (0.930)	0.224 (0.139)	.	.	0.188
<i>FF3</i>	-0.298 (1.245)	1.365 (1.292)	.	0.629 (0.529)	1.063 (0.505)	0.891

Conclusions

- There is significant predictability in consumption growth! The expected growth co-varies with the business cycles.
- A VAR specification identifies two distinct types of growth shocks. The expected growth shock is a significantly priced risk factor in the financial market.
- Find a risk-based explanation for size and B/M effects in terms of exposures to expected growth news.
- Macroeconomic news related to business cycles serve as financial market risk factors in a manner that is consistent with their predictability to aggregate growth.
- Aggregate cash-flow dynamics is important for intertemporal asset pricing.

Future Work

- Individual asset cash flow dynamics and its implications for the asset risk premium. (particularly interesting for the momentum effect).
- Allow time-varying volatilities and hence time-varying risk premia and volatility risk (Tauchen (2004)).
- Mechanisms for hedging macroeconomic and aggregate consumption risks.